### Introduction to Bayesian Statistics (II)

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BDSI 2022

### THE BAYESIAN MACHINERY

1. Set up parametric models





2. Compute posterior distribution using Bayes' theorem

#### EXAMPLE: BIOASSAY EXPERIMENT

The scientific problem:

- Bioassay experiment: toxicity tests on animals
- Various dose levels of the drug compounds apply to batches of animals
- Responses typically characterized by a binary outcome: alive or dead

The data

- of the form  $(x_i, n_i, y_i), i = 1, ..., k$
- x<sub>i</sub>: dosage of the *i*th dose level
- *n<sub>i</sub>*: number of animals experimented with *i*th dose level
- *y<sub>i</sub>*: number of positive outcomes for the *i*th dose level (*y<sub>i</sub>* ≤ *n*)

### SAMPLE DATA: BIOASSAY EXPERIMENT

Dose, $x_i$ (log g/ml)	Number of animals, $n_i$	Number of deaths, $y_i$
-0.86	5	0
-0.30	5	1
-0.05	5	3
0.73	5	5

MODELING DOSE-RESPONSE RELATIONSHIP

Likelihood modeling

 Outcomes of animals within each dosage group are exchangeable: reasonable to assume

 $y_i \mid \theta_i \sim \operatorname{Bin}(n_i, \theta_i)$ 

- Linking  $\theta_i$  to dose levels
  - Linear relation?

$$\theta_i = \alpha + \beta x_i$$

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Logistic modeling

$$\theta_i = \frac{\exp(\alpha + \beta x_i)}{1 + \exp(\alpha + \beta x_i)}$$

or equivalently,

 $logit(\theta_i) = \alpha + \beta x_i$ 

### THE PRIOR DISTRIBUTION

specify parametric model for  $p(\alpha, \beta)$ 

Without any prior knowledge, a flat prior is often assumed

 $p(\alpha,\beta) \propto 1$ 

• Implies  $\alpha$  and  $\beta$  are independent

• All combinations of  $(\alpha, \beta)$  are equally likely

Also an example of improper prior

Is the prior informative?

### **APPLYING BAYES' THEOREM**

The likelihood function

$$p(y_i \mid \alpha, \beta, n_i, x_i) \propto \left[\frac{\exp(\alpha + \beta x_i)}{1 + \exp(\alpha + \beta x_i)}\right]^{y_i} \left[1 - \frac{\exp(\alpha + \beta x_i)}{1 + \exp(\alpha + \beta x_i)}\right]^{n-y_i}$$



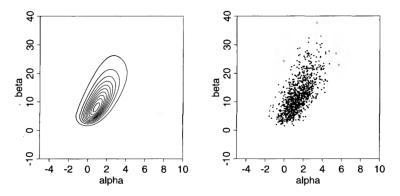
 $p(\alpha,\beta)\propto 1$ 

The posterior

$$p(\alpha,\beta \mid (x_1,n_1,y_1),...,(x_k,n_k,y_k)) \propto \prod_{i=1}^k p(y_i \mid \alpha,\beta,n_i,x_i)$$

Even for such simple problem, there is no closed-form expression for the posterior distribution

#### **R**EPRESENTING THE POSTERIOR DISTRIBUTION



Left panel: contour plot for the posterior density

 Right panel: scatter plot of 1000 draws from the posterior distribution

#### **R**EMARKS ON THE BIOASSAY EXPERIMENT

The posterior indicates large degree of uncertainty on (α, β)

 Posterior mode coincides with the maximum likelihood estimates (why?)

What is the impact of our prior choice?

Prior is part of the model assumptions

"All models are wrong, but some are useful"

In general, priors have less impact on the posteriors if the data are strongly informative

## COMMON CONSIDERATIONS FOR PRIOR CONSTRUCTION

▶ Enable analytic computation of posteriors: conjugate prior

- Example: normal prior for the mean parameter of a normal likelihood
- Jefferey's invariance principle
- Flat/"non-informative" prior

Use of informative prior from expert's opinion or existing data

#### **BAYESIAN COMPUTATION**

- Posterior computation via Bayes' theorem is typically intractable except a few special cases (e.g., conjugate priors)
- In most cases, priors and likelihood functions can be analytically computed
- The normalizing constant, which requires computing an integral, is difficult to obtain
- The general strategy for posterior computation is by numerical approximation: Markov Chain Monte Carlo (MCMC) algorithms

### BAYESIAN MODEL FOR VARIABLE SELECTION

- The problem: Identify potentially multiple associations from a linear model
- The likelihood is based on a multiple linear regression model

$$\boldsymbol{y} = \mu \mathbf{1} + \sum_{j=1}^{p} \beta_j \boldsymbol{g}_j + \boldsymbol{e}, \ \boldsymbol{e} \sim N(\mathbf{0}, \tau^{-1} \boldsymbol{I})$$

- Define γ<sub>j</sub> := 1(β<sub>j</sub> ≠ 0), a latent indicator of association status for the *j*th predictor
- Interested in infer  $\gamma := (\gamma_1, ..., \gamma_p)$
- In statistical terms,  $\beta_j$ 's and  $\tau$  are considered nuisance parameters.

# BAYESIAN MODEL FOR VARIABLE SELECTION (CONT'D)

Priors

• 
$$\gamma_i$$
's are assumed exchangeable

 $\gamma_i \mid \theta \sim \text{Bernoulli}(\theta)$ 

$$\triangleright \ \beta_j \mid \gamma_j = 1 \ \sim \ \mathbf{N}(0, 1)$$

The induced marginal prior on β<sub>j</sub> is known as a "spike-and-slab" prior



# BAYESIAN MODEL FOR VARIABLE SELECTION (CONT'D)

Considerations for hyper-parameters  $\theta$ , a, b

- ► In most cases, we believe true association signals are sparse and set  $\theta = \frac{1}{p}$
- For reasonably large sample size, the inference result is less sensitive to the choice of (*a*, *b*)

• Typically set a flat prior by letting  $a \rightarrow 0, b \rightarrow 0$ .

## INFERENCE FOR BAYESIAN VARIABLE SELECTION MODEL

- Need to find  $Pr(\gamma \mid y, G)$
- Compute marginal likelihood function  $P(y \mid \gamma, G)$ 
  - The computation requires to integrate out the nuisance parameters β<sub>k</sub>'s and τ
  - In this case, the marginal likelihood can be analytically evaluated
- What is the normalizing constant?

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$$\sum_{\boldsymbol{\gamma}'} \Pr(\boldsymbol{\gamma}') P(\boldsymbol{y} \mid \boldsymbol{\gamma}', \boldsymbol{G})$$

For large *p*, the exact computation is practically impossible, approximated computation is warranted.

## ADVANTAGES AND CHALLENGES IN BAYESIAN VARIABLE SELECTION MODEL

Advantages

- It is a model designed for discovering associations (not for predictions)
- It is expandable if annotation data (*d<sub>j</sub>*'s) on predictors become available

$$\operatorname{logit}(\theta_j) = \alpha + \beta d_j$$

The posterior provides comprehensive information rather than reporting a single "best" model

Challenges

computation and scalability

#### FUTURE DIRECTIONS

#### Recommended readings



Topics not thoroughly discussed

- Bayesian computation
- Bayesian model diagnosis
- Bayesian Frequentist compromise